

## LECTURE - (14)

### Agenda:

- ① Consistency of a statistical estimator
- ② Examples
- ③ LEFTOVER FROM CHAPTER 8: Confidence interval for  $\sigma^2$ .

### CONSISTENCY OF A STATISTICAL ESTIMATOR

Let  $\hat{\theta}$  be a statistical estimator of a parameter  $\theta$  obtained from a sample of size  $n$ .

Definition:  $\hat{\theta}$  is said to be a consistent estimator of  $\theta$  if, for any positive number  $\epsilon$ ,

$$\lim_{n \rightarrow \infty} P(|\hat{\theta}_n - \theta| \leq \epsilon) = 1.$$

OR  $\hat{\theta}$  converges to  $\theta$  in probability.

Intuitively, what this means is that if the sample size is large enough, the probability that the estimator  $\hat{\theta}$  is very close to the true value  $\theta$  is almost 1. This is a very desirable quality in a statistical estimator. Any estimator that is not

consistent should be viewed suspiciously.

How do we prove that an estimator is consistent?  
Here are 2 results which are extremely useful.

RESULT 1: If  $MSE(\hat{\theta}) \rightarrow 0$  as  $n \rightarrow \infty$ , then  $\hat{\theta}$  is a consistent estimator of  $\theta$ .

RESULT 2: Let  $\hat{\theta}$  be a consistent estimator of  $\theta$ , and  $\hat{\theta}'$  be a consistent estimator of  $\theta'$ . Then,

(i)  $\hat{\theta} + \hat{\theta}'$  is a consistent estimator of  $\theta + \theta'$ .

(ii)  $\hat{\theta} \times \hat{\theta}'$  is a consistent estimator of  $\theta \times \theta'$ .

(iii) If  $\theta' \neq 0$ , then  $\frac{\hat{\theta}}{\hat{\theta}'}$  is a consistent estimator

of  $\frac{\theta}{\theta'}$ .

(iv) If  $g$  is a real-valued function that is continuous at  $\theta$ , then  $g(\hat{\theta}_n)$  is a consistent estimator for  $g(\theta)$ .

Let us look at a few examples.

Example 1: Let  $Y_1, Y_2, \dots, Y_n$  be a random sample from a population with mean  $\mu$  and finite variance  $\sigma^2$ . Is  $\bar{Y}$  consistent for  $\mu$ ?

Since  $MSE(\bar{Y}) = \frac{\sigma^2}{n} \rightarrow 0$  as  $n \rightarrow \infty$ ,

$\bar{Y}$  is in fact a consistent estimator of  $\mu$ . We noted

this fact in STA4321 as the weak law of large numbers.

Example 2: Let  $Y_1, Y_2, \dots, Y_n$  be a random sample from a population ~~which~~ which is Normal( $\mu, \sigma^2$ ). Is  $S^2$  a consistent estimator of  $\sigma^2$ ?

Since  $MSE(S^2) = \frac{2\sigma^4}{n-1} \rightarrow 0$  as  $n \rightarrow \infty$ ,

$S^2$  is in fact a consistent estimator of  $\sigma^2$ .

Example 3: Let  $Y_1, Y_2, \dots, Y_n$  be a random sample from a population which is Normal( $\mu, \sigma^2$ ). Provide a consistent estimator for  $\frac{\mu}{\sigma}$ .

Note that,

$\bar{Y}$  is consistent for  $\mu$ .

$S^2$  is consistent for  $\sigma^2 \Rightarrow \sqrt{S^2} = S$  is consistent for  $\sqrt{\sigma^2} = \sigma$  (as  $g(x) = \sqrt{x}$  is a ~~continuous~~ continuous function ~~for~~ for  $x > 0$ ).

Hence,  $\frac{\bar{Y}}{S}$  is a consistent estimator for  $\frac{\mu}{\sigma}$ .

LEFTOVER FROM CHAPTER 8: CONFIDENCE INTERVAL FOR  $\sigma^2$ .

Suppose we have a random sample  $Y_1, Y_2, \dots, Y_n$  from a population which is Normal  $(\mu, \sigma^2)$ ,  $\mu$  and  $\sigma^2$  are both unknown. The goal is to construct a confidence interval for  $\sigma^2$  with confidence level  $1 - \alpha$ .

Step 1: Note that in Exercise 8.41, we saw that if  $\mu = 0$ , then  $\frac{\sum_{i=1}^n Y_i^2}{\sigma^2}$  has a  $\chi^2$ -distribution

with  $n$  degrees of freedom, and hence can be used as a pivotal quantity (The exercise only deals with  $n=1$ ), but the result holds for general  $n$ ). However, in the current problem,  $\mu$  is unknown. But we know that

$\frac{(n-1)S^2}{\sigma^2}$  has a  $\chi^2$ -distribution with

$(n-1)$  degrees of freedom. Hence,

~~$\frac{(n-1)S^2}{\sigma^2}$~~   $\frac{(n-1)S^2}{\sigma^2}$  is a pivotal quantity.

Step 2: Find  $a$  and  $b$  such that

$$P\left(a \leq \frac{(n-1)S^2}{\sigma^2} \leq b\right) = 1-\alpha.$$

Note that ~~the  $(1-\alpha)$ th quantile of the  $\chi^2$  distribution with  $n$  degrees of freedom.~~  $\chi^2_{n, 1-\alpha}$  denotes

the  $(1-\alpha)$ <sup>th</sup> quantile of the  $\chi^2$ -distribution with  $n$  degrees of freedom.

FACT:  $P\left(\chi^2_{n-1, \frac{\alpha}{2}} \leq \frac{(n-1)S^2}{\sigma^2} \leq \chi^2_{n-1, 1-\frac{\alpha}{2}}\right) = 1-\alpha.$

Hence,  $a = \chi^2_{n-1, \frac{\alpha}{2}}$  and  $b = \chi^2_{n-1, 1-\frac{\alpha}{2}}$ .

Step 3: Express  $a \leq \frac{(n-1)S^2}{\sigma^2} \leq b$  as  $\hat{\sigma}_L^2 \leq \sigma^2 \leq \hat{\sigma}_U^2$ .

Note that

$$a \leq \frac{(n-1)S^2}{\sigma^2} \leq b$$

$$\Leftrightarrow \frac{(n-1)S^2}{\chi^2_{n-1, 1-\frac{\alpha}{2}}} \leq \sigma^2 \leq \frac{(n-1)S^2}{\chi^2_{n-1, \frac{\alpha}{2}}}.$$

Hence, a  $(1-\alpha)$  confidence interval for  $\sigma^2$  is given by

$$\left[ \frac{(n-1)S^2}{\chi^2_{n-1, 1-\frac{\alpha}{2}}}, \frac{(n-1)S^2}{\chi^2_{n-1, \frac{\alpha}{2}}} \right].$$