

LECTURE - (26)

Agenda:

- Small-sample hypothesis testing for μ and $\mu_1 - \mu_2$.

Consider the following situation. Suppose Y_1, Y_2, \dots, Y_n is a random sample from a population which is normal with mean μ . We wish to test the hypotheses

$$H_0: \mu = \mu_0 \quad \text{v.s.} \quad H_A: \mu > \mu_0.$$

with level α . HOWEVER, THE SAMPLE SIZE n IS NOT LARGE ENOUGH TO APPROXIMATE THE DISTRIBUTION OF $\frac{\sqrt{n}(\bar{Y} - \mu)}{S}$ BY THE NORMAL $(0, 1)$ DISTRIBUTION.

But this does not pose a problem, since the assumption that the population is normal implies that

$\frac{\sqrt{n}(\bar{Y} - \mu)}{S}$ has a t -distribution with

$n-1$ degrees of freedom.

Recall that $t_{n-1, 1-\alpha}$ denotes the $(1-\alpha)^{\text{th}}$ quantile of the t -distribution with $n-1$ degrees of freedom.

RESULT: If Y_1, Y_2, \dots, Y_n is Normal (μ, σ^2) with μ, σ^2 both unknown, then the level- α test for testing $H_0: \mu = \mu_0$ v.s. $H_A: \mu > \mu_0$ is given by

TEST STATISTIC: \bar{Y}

REJECTION REGION: $\bar{Y} > \mu_0 + \frac{t_{n-1, 1-\alpha} S}{\sqrt{n}}$

This testing procedure is known as the exact or small-sample testing procedure for μ .

RESULT: If Y_1, Y_2, \dots, Y_n is Normal (μ, σ^2) with μ, σ^2 both unknown, then the level- α test for testing $H_0: \mu = \mu_0$ v.s. $H_A: \mu < \mu_0$ is given by

TEST STATISTIC: \bar{Y}

REJECTION REGION: $\bar{Y} < \mu_0 + \frac{t_{n-1, \alpha} S}{\sqrt{n}}$

RESULT: If Y_1, Y_2, \dots, Y_n is Normal (μ, σ^2) with μ, σ^2 both unknown, then the level- α test for testing $H_0: \mu = \mu_0$ v.s. $H_A: \mu \neq \mu_0$ is given by

TEST STATISTIC: \bar{Y}

REJECTION REGION: $\bar{Y} < \mu_0 + \frac{t_{n-1, \frac{\alpha}{2}} S}{\sqrt{n}}$ or

$\bar{Y} > \mu_0 + \frac{t_{n-1, 1-\frac{\alpha}{2}} S}{\sqrt{n}}$

As you can see, the only difference between the large sample test and exact test for μ is that the Normal z quantiles are replaced by the corresponding t quantiles.

Consider the following situation. Suppose we have random samples from two normal populations:

$Y_{11}, Y_{12}, \dots, Y_{1n_1}$ is a random sample from a Normal (μ_1, σ^2) population, and $Y_{21}, Y_{22}, \dots, Y_{2n_2}$ is a random sample from a Normal (μ_2, σ^2) population.

We want to test three different sets of hypotheses.

- (1) ~~.....~~ $H_0: \mu_1 - \mu_2 = D_0$ v.s. $H_A: \mu_1 - \mu_2 \neq D_0$.
- (2) $H_0: \mu_1 - \mu_2 = D_0$ v.s. $H_A: \mu_1 - \mu_2 < D_0$.
- (3) $H_0: \mu_1 - \mu_2 = D_0$ v.s. $H_A: \mu_1 - \mu_2 \neq D_0$.

Recall that $\frac{(\bar{Y}_1 - \bar{Y}_2) - (\mu_1 - \mu_2)}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$ has a

t-distribution with $n_1 + n_2 - 2$ degrees of freedom, where $S_p^2 = \frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2}$ is the pooled sample variance.

RESULT: Suppose $Y_{11}, Y_{12}, \dots, Y_{1n_1}$ and $Y_{21}, Y_{22}, \dots, Y_{2n_2}$ are independent random samples from a Normal (μ_1, σ^2) and Normal (μ_2, σ^2) population respectively.

(1) The level- α testing procedure for testing $H_0: \mu_1 - \mu_2 = D_0$ v.s. $H_A: \mu_1 - \mu_2 > D_0$ is given by

TEST STATISTIC: $\bar{Y}_1 - \bar{Y}_2$

REJECTION REGION:

$$\bar{Y}_1 - \bar{Y}_2 > D_0 + t_{n_1 + n_2 - 2, 1 - \alpha} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

(2) The level- α testing procedure for testing $H_0: \mu_1 - \mu_2 = D_0$ v.s. $H_A: \mu_1 - \mu_2 < D_0$ is given by

TEST STATISTIC: $\bar{Y}_1 - \bar{Y}_2$

REJECTION REGION:

$$\bar{Y}_1 - \bar{Y}_2 < D_0 + t_{n_1+n_2-2, \alpha} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

- ③ The level- α testing procedure for testing $H_0: \mu_1 - \mu_2 = D_0$ v.s. $H_A: \mu_1 - \mu_2 \neq D_0$ is given by

TEST STATISTIC: $\bar{Y}_1 - \bar{Y}_2$

REJECTION REGION:

$$\bar{Y}_1 - \bar{Y}_2 < D_0 + t_{n_1+n_2-2, \frac{\alpha}{2}} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}} \quad \text{or}$$

$$\bar{Y}_1 - \bar{Y}_2 > D_0 + t_{n_1+n_2-2, 1-\frac{\alpha}{2}} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

