

LECTURE- (9)

Agenda:

- ① Confidence interval in normal population with UNKNOWN σ .
- ② Upper confidence limit.
- ③ Lower confidence limit.
- ④ Examples

CONFIDENCE INTERVAL IN NORMAL POPULATION WITH UNKNOWN σ .

In the last lecture we considered the scenario, where we have a random sample Y_1, Y_2, \dots, Y_n from a population which is Normal (μ, σ^2) . We are interested in a 95% confidence interval for μ . If σ is known, we used $\frac{\sqrt{n}(\bar{Y} - \mu)}{\sigma}$ as the pivotal quantity to obtain the

$$\text{confidence interval } \left[\bar{Y} - \frac{1.96\sigma}{\sqrt{n}}, \bar{Y} + \frac{1.96\sigma}{\sqrt{n}} \right]$$

WHAT IF σ is UNKNOWN? Let $s^2 = \frac{1}{n-2} \sum_{i=1}^n (Y_i - \bar{Y})^2$.

- ① Recall that $\frac{\sqrt{n}(\bar{Y} - \mu)}{s}$ has a t-distribution with $n-2$ degrees of freedom. Since this distribution is independent of μ , $g(\text{Data}, \mu) = \frac{\sqrt{n}(\bar{Y} - \mu)}{s}$ is a pivotal quantity.

(2) We want a and b such that

$$P\left(a \leq \frac{\sqrt{n}(\bar{Y} - \mu)}{S} \leq b\right) = 95\%.$$

FACT: Let $t_{n-1, 1-\alpha}$ denote the $(1-\alpha)^{\text{th}}$ quantile for every $0 \leq \alpha \leq 1$ of the t -distribution with $n-1$ degrees of freedom.

By the symmetry of the t -distribution around 0, it is known that for every $0 \leq \alpha \leq 1$,

$$P\left(-t_{n-1, \frac{1-\alpha}{2}} \leq \frac{\sqrt{n}(\bar{Y} - \mu)}{S} \leq t_{n-1, \frac{1-\alpha}{2}}\right) = 1-\alpha.$$

We want $1-\alpha = 95\%$ in this particular problem.

Hence $1 - \frac{\alpha}{2} = 97.5\%$. We get that

$$P\left(-t_{n-1, 97.5\%} \leq \frac{\sqrt{n}(\bar{Y} - \mu)}{S} \leq t_{n-1, 97.5\%}\right) = 95\%.$$

$$(3) \quad -t_{n-1, 97.5\%} \leq \frac{\sqrt{n}(\bar{Y} - \mu)}{S} \leq t_{n-1, 97.5\%}.$$

$$\Leftrightarrow \frac{-t_{n-1, 97.5\%} S}{\sqrt{n}} \leq \bar{Y} - \mu \leq \frac{t_{n-1, 97.5\%} S}{\sqrt{n}}$$

$$\Leftrightarrow \bar{Y} - \frac{t_{n-1, 97.5\%} S}{\sqrt{n}} \leq \mu \leq \bar{Y} + \frac{t_{n-1, 97.5\%} S}{\sqrt{n}}$$

Hence, a ~~95%~~ 95% confidence interval for μ is given by $\left[\bar{Y} - \frac{t_{n-1, 97.5\%} S}{\sqrt{n}}, \bar{Y} + \frac{t_{n-1, 97.5\%} S}{\sqrt{n}} \right]$

UPPER CONFIDENCE LIMIT ~~IS~~

DEFINITION: We say that $\hat{\theta}_{upper}$ is an upper confidence limit for an unknown parameter θ ~~with~~ with confidence level $1-\alpha$, if

$$P(\theta \leq \hat{\theta}_{upper}) = 1-\alpha.$$

Method to find upper confidence limit with confidence level $1-\alpha$

- ① Get a pivotal quantity $g(\text{Data}, \theta)$ WHICH IS AN INCREASING FUNCTION OF θ .
- ② Find the ~~critical~~ $(1-\alpha)^{th}$ quantile of $g(\text{Data}, \theta)$, i.e., find b such that

$$P(g(\text{Data}, \theta) \leq b) = 1-\alpha.$$

- ③ Reduce $g(\text{Data}, \theta) \leq b$ to the form $\theta \leq \hat{\theta}_{upper}$.

Let us try this method to find an upper confidence limit for μ in the normal population problem with KNOWN σ .

- ① The pivotal quantity we generally work with is $\frac{\sqrt{n}(\bar{Y} - \mu)}{\sigma}$. However since we want the pivotal quantity to be an increasing function of μ , we just

choose the negative, i.e., $\frac{\sqrt{n}(\mu - \bar{Y})}{\sigma}$ as our

pivotal quantity. Suppose the confidence level that we desire is $1 - \alpha$.

(2) We want b such that

$$P\left(\frac{\sqrt{n}(\mu - \bar{Y})}{\sigma} \leq b\right) = 1 - \alpha$$

$$\Leftrightarrow P\left(-\frac{\sqrt{n}(\mu - \bar{Y})}{\sigma} \geq -b\right) = 1 - \alpha$$

$$\Leftrightarrow P\left(\frac{\sqrt{n}(\bar{Y} - \mu)}{\sigma} \geq -b\right) = 1 - \alpha$$

$$\Leftrightarrow P\left(\frac{\sqrt{n}(\bar{Y} - \mu)}{\sigma} \leq -b\right) = \alpha$$

Hence, $-b$ is the α th quantile of $Z = \frac{\sqrt{n}(\bar{Y} - \mu)}{\sigma}$,

which we know has a Normal(0, 1) distribution.

The standard notation for the α th quantile of the Normal(0, 1) distribution is z_α for every $0 \leq \alpha \leq 1$.

Hence,

$$-b = z_\alpha \Rightarrow b = -z_\alpha = z_{1-\alpha}$$

FACT: By the symmetry of the Normal(0, 1) distribution.

$$(3) \quad \frac{\sqrt{n}(\mu - \bar{Y})}{\sigma} \leq b \Leftrightarrow \mu \leq \bar{Y} + \frac{b\sigma}{\sqrt{n}} = \bar{Y} + \frac{z_{1-\alpha}\sigma}{\sqrt{n}}$$

Hence, a $(1-\alpha)$ upper confidence limit for μ is given by $\hat{\mu}_{upper} = \bar{Y} + \frac{Z_{1-\alpha} \sigma}{\sqrt{n}}$.

LOWER CONFIDENCE LIMIT

DEFINITION: We say that $\hat{\theta}_{lower}$ is a lower confidence limit for an unknown parameter θ with confidence level $1-\alpha$, if

$$P(\hat{\theta}_{lower} \leq \theta) = 1-\alpha$$

Method to find lower confidence limit with confidence level $1-\alpha$

- (1) Get a pivotal quantity $g(\text{Data}, \theta)$ WHICH IS AN INCREASING FUNCTION OF θ .
- (2) Find b such that

$$P(b \leq g(\text{Data}, \theta)) = 1-\alpha$$

- (2) Reduce $b \leq g(\text{Data}, \theta)$ to the form $\hat{\theta}_{lower} \leq \theta$.

Let us try this method to find a lower confidence limit for μ in the normal population problem with KNOWN σ .

- (1) Again, we use the pivotal quantity $\frac{\sqrt{n}(\mu - \bar{Y})}{\sigma}$ to make sure it is an increasing function of μ .

(2) We want to find b such that

$$P\left(b \leq \frac{\sqrt{n}(\mu - \bar{Y})}{\sigma}\right) = 1 - \alpha$$

$$\Leftrightarrow P\left(\frac{\sqrt{n}(\bar{Y} - \mu)}{\sigma} \leq -b\right) = 1 - \alpha \rightarrow \text{(Normal}(0,1)\text{)}$$

$$\Leftrightarrow -b = z_{1-\alpha} \quad (\text{the } (1-\alpha)^{\text{th}} \text{ quantile of the Normal}(0,1) \text{ distribution})$$

$$\Leftrightarrow b = -z_{1-\alpha}$$

$$(3) \quad b \leq \frac{\sqrt{n}(\mu - \bar{Y})}{\sigma} \Leftrightarrow \bar{Y} + \frac{b\sigma}{\sqrt{n}} \leq \mu \Leftrightarrow \bar{Y} - \frac{z_{1-\alpha}\sigma}{\sqrt{n}} \leq \mu.$$

Hence, a $(1-\alpha)$ lower confidence limit for μ is given by

$$\hat{\mu}_{\text{lower}} = \bar{Y} - \frac{z_{1-\alpha}\sigma}{\sqrt{n}}$$