

$$\underline{6.26} \quad \boxed{(a)} \quad P(X < 2, Y > 1) = P((X, Y) \in A),$$

$$\text{where } A = \{(x, y) : x < 2, y > 1\}.$$

Range of density  $f_{X,Y}$  is denoted by  $R$ .

$$R = \{(x, y) : 0 \leq y \leq x < \infty\}$$

$$\begin{aligned} \Rightarrow A \cap R &= \{(x, y) : x < 2, y > 1, 0 \leq y \leq x < \infty\} \\ &= \{(x, y) : 1 \leq x < 2, 1 \leq y \leq x\} \end{aligned}$$

Hence,

$$\begin{aligned} P((X, Y) \in A) &= \iint_{A \cap R} f_{X,Y}(x, y) \, dy \, dx \\ &= \int_1^2 \int_1^x e^{-x} \, dy \, dx \\ &= \int_1^2 (x-1) e^{-x} \, dx \\ &= \int_1^2 x e^{-x} \, dx - \int_1^2 e^{-x} \, dx \\ &= \left[ -x e^{-x} \right]_1^2 + \int_1^2 e^{-x} \, dx - \int_1^2 e^{-x} \, dx \\ &= e^{-1} - 2e^{-2} \end{aligned}$$

$$\boxed{(b)} \quad P(X \geq 2Y) = P((X, Y) \in A), \text{ where } A = \{(x, y) : x \geq 2y\}$$

$$\Rightarrow A \cap R = \{(x, y) : x \geq 2y, 0 \leq y \leq x < \infty\}$$

$$= \left\{ (x, y) : 0 \leq x < \infty, 0 \leq y \leq \frac{x}{2} \right\}$$

Hence,  $P((X, Y) \in A)$

$$= \iint_{A \cap R} f_{X, Y}(x, y) dy dx$$

$$= \int_0^{\infty} \int_0^{x/2} e^{-x} dy dx$$

$$= \int_0^{\infty} \frac{x}{2} e^{-x} dx$$

$$= \frac{\Gamma(2)}{2}$$

$$= \frac{1}{2} .$$

(c)

$$f_X(x) = \int_{-\infty}^{\infty} f_{X, Y}(x, y) dy$$

$$= \begin{cases} \int_0^x e^{-x} dy & \text{if } x \geq 0, \\ 0 & \text{if } x < 0. \end{cases}$$

$$= \begin{cases} x e^{-x} & \text{if } x \geq 0, \\ 0 & \text{if } x < 0. \end{cases}$$

$$\boxed{(d)} \quad f_{X|Y=y}(x) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$$

$$\text{Note that } f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx$$

$$= \begin{cases} \int_y^{\infty} e^{-x} dx & \text{if } y \geq 0, \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} e^{-y} & \text{if } y \geq 0, \\ 0 & \text{otherwise} \end{cases}$$

$$\text{Hence, } f_{X|Y=y}(x) = \begin{cases} \frac{e^{-x}}{e^{-y}} & \text{if } y \leq x < \infty, \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} e^{-(x-y)} & \text{if } y \leq x < \infty, \\ 0 & \text{otherwise} \end{cases}$$

**6.68**

a. For the next five cars entering this intersections, let  $X_1$  denote the number of cars that turn left,  $X_2$  denote the number of cars that turn right,  $X_3$  denote the number of cars that continue straight ahead, then  $X_1, X_2, X_3$  has the multinomial distribution with  $p_1 = 0.4, p_2 = 0.25, p_3 = 0.35$ , then the probability of interest is

$$P(X_1 = 1, X_2 = 1, X_3 = 3) = \frac{5!}{1!1!3!} 0.4 \cdot 0.25 \cdot 0.35^3 = 0.08575$$

b.  $X_2$  has the binomial distribution with  $n = 5, p_2 = 0.25$ , then

$$\begin{aligned} P(X_2 \geq 1) &= 1 - P(X_2 = 0) \\ &= 1 - \binom{5}{0} 0.25^0 (1 - 0.25)^5 \\ &= 0.7627 \end{aligned}$$

c. For the 100 cars entering this intersections, let  $Y$  denote the number of cars that turn left,  $Y$  has the binomial distribution with  $p_1 = 0.4$ , then

$$\begin{aligned} E(Y) &= np_1 = 40 \\ V(Y) &= np_1(1 - p_1) = 24 \end{aligned}$$

The assumption is that the trials are independent.

**6.70**

a. Let  $X_1 \sim N(\mu_1, \sigma_1^2), X_2 \sim N(\mu_2, \sigma_2^2)$ , then the moment generating function of

$Y = aX_1 + bX_2$  is

$$\begin{aligned} M_Y(t) &= E(e^{tY}) \\ &= E(e^{taX_1} e^{tbX_2}) \\ &= E(e^{taX_1}) E(e^{tbX_2}) \\ &= M_{X_1}(at) M_{X_2}(bt) \\ &= e^{\mu_1 at + \sigma_1^2 a^2 t^2 / 2} e^{\mu_2 bt + \sigma_2^2 b^2 t^2 / 2} \\ &= e^{(a\mu_1 + b\mu_2)t + (a^2\sigma_1^2 + b^2\sigma_2^2)t^2 / 2} \end{aligned}$$

This is the moment-generating function of a normal distribution,

thus  $Y \sim N(a\mu_1 + b\mu_2, a^2\sigma_1^2 + b^2\sigma_2^2)$

b.

$$E(Y) = \frac{d}{dt} M_Y(t) \Big|_{t=0} = \left[ (a\mu_1 + b\mu_2) + (a^2\sigma_1^2 + b^2\sigma_2^2)t \right] e^{(a\mu_1 + b\mu_2)t + (a^2\sigma_1^2 + b^2\sigma_2^2)t^2/2} \Big|_{t=0}$$

$$= a\mu_1 + b\mu_2$$

$$V(Y) = \frac{d^2}{dt^2} M_Y(t) \Big|_{t=0} = \left[ (a^2\sigma_1^2 + b^2\sigma_2^2) + \left( (a\mu_1 + b\mu_2) + (a^2\sigma_1^2 + b^2\sigma_2^2)t \right)^2 \right] e^{(a\mu_1 + b\mu_2)t + (a^2\sigma_1^2 + b^2\sigma_2^2)t^2/2} \Big|_{t=0}$$

$$= a^2\sigma_1^2 + b^2\sigma_2^2$$

### 6.86

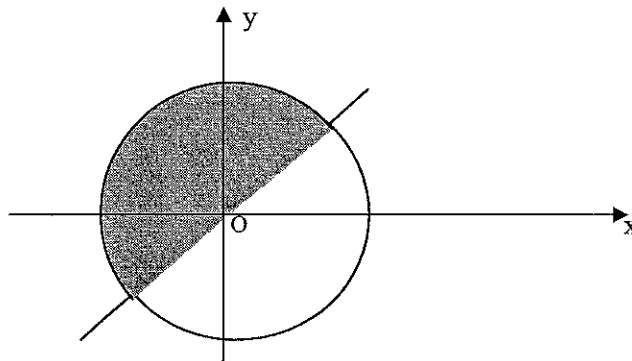
a. for a fixed  $X = x \in [-1, 1]$ , we have  $-\sqrt{1-x^2} \leq y \leq \sqrt{1-x^2}$ , the marginal density function of  $X$  is

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$= \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{1}{\pi} dy$$

$$= \frac{2}{\pi} \sqrt{1-x^2}, \quad -1 \leq x \leq 1$$

b. define the set  $A = \{(x, y) : x^2 + y^2 \leq 1, x \leq y\}$ , then  $A$  is the shadow area in the figure below.



The probability of interest is

$$\begin{aligned}
P(X \leq Y) &= \iint_A f(x, y) dx dy \\
&= \frac{1}{\pi} S(A) \\
&= \frac{1}{\pi} \cdot \frac{\pi}{2} \\
&= \frac{1}{2}
\end{aligned}$$

### 7.3

a.  $X$  have the negative binomial distribution with  $r = 4, p = 0.4$ , then the distribution of  $X$  is

$$\begin{aligned}
P(X = x) &= \binom{x+r-1}{r-1} p^r (1-p)^x \\
&= \binom{x+3}{3} 0.4^4 \cdot 0.6^x, \quad x = 0, 1, 2, \dots
\end{aligned}$$

b.  $Y = X + 4$

c. the distribution of  $Y$  is

$$\begin{aligned}
P(Y = y) &= P(X + 4 = y) \\
&= P(X = y - 4) \\
&= \binom{y-4+3}{3} 0.4^4 \cdot 0.6^{y-4} \\
&= \binom{y-1}{3} 0.4^4 \cdot 0.6^{y-4}, \quad y = 4, 5, 6, \dots
\end{aligned}$$

### 7.9

For  $y \in (1, e)$ ,

$$\begin{aligned}
F_Y(y) &= P(Y < y) \\
&= P(e^X < y) \\
&= P(X < \ln y) \\
&= \int_0^{\ln y} 1 dx \\
&= \ln y
\end{aligned}$$

When  $y \leq 1, F_Y(y) = 0$ , when  $y \geq e, F_Y(y) = 1$

Thus the probability density function of  $Y$  is

$$f_Y(y) = \frac{d}{dx} F_Y(y) = \begin{cases} \frac{1}{y}, & 1 < y < e \\ 0, & \text{otherwise} \end{cases}$$

### 7.11

For  $y \geq 0$ ,

$$\begin{aligned} F_Y(y) &= P(Y < y) \\ &= P(cX < y) \\ &= P(X < \frac{1}{c}y) \\ &= \int_0^{\frac{1}{c}y} \frac{1}{\theta} e^{-\frac{1}{\theta}x} dx \\ &= 1 - e^{-\frac{1}{c\theta}y} \end{aligned}$$

Then the probability density function of  $Y$  is

$$f_Y(y) = \frac{d}{dx} F_Y(y) = \begin{cases} \frac{1}{c\theta} e^{-\frac{1}{c\theta}y}, & y \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

Thus  $Y \sim \text{exponential}(c\theta)$

### 7.13

For  $y > 0$ ,

$$\begin{aligned} F_Y(y) &= P(Y < y) \\ &= P(e^X < y) \\ &= P(X < \ln y) \\ &= F_X(\ln y) \end{aligned}$$

Differentiation this equation, we have

$$\begin{aligned} f_Y(y) &= \frac{1}{y} f_X(\ln y) \\ &= \frac{1}{y} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(\ln y - \mu)^2}{2\sigma^2}}, \quad y > 0 \end{aligned}$$

Otherwise,  $f_Y(y) = 0$

7.20

$$Y = e^X = g(X) \quad f_X(x) = \begin{cases} 1 & \text{if } x \in (0, 1), \\ 0 & \text{otherwise} \end{cases}$$

Step 1:  $(\alpha, \beta) = (e^0, e^1) = (1, e)$ .

Step 2:  $Y = e^X \Rightarrow X = \log Y = h(Y)$ .

Step 3:  $h'(y) = \frac{1}{y} \neq 0$  if  $y \in (1, e)$ .

Step 4:

$$f_Y(y) = \begin{cases} f_X(h(y)) |h'(y)| & \text{if } y \in (1, e), \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} \frac{1}{y} & \text{if } y \in (1, e), \\ 0 & \text{otherwise} \end{cases}$$

7.27

$$f_{X_1, X_2}(x_1, x_2) = \begin{cases} 2 & \text{if } 0 \leq x_1 \leq 1, 0 \leq x_2 \leq 1, 0 \leq x_1 + x_2 \leq 1, \\ 0 & \text{otherwise.} \end{cases}$$

$$Y = X_1 + X_2$$

$$f_{X_2}(x_2) = \begin{cases} \int_0^{1-x_2} 2 dx_1 & \text{if } 0 \leq x_2 \leq 1, \\ 0 & \text{otherwise} \end{cases} = \begin{cases} 2(1-x_2) & \text{if } 0 \leq x_2 \leq 1, \\ 0 & \text{otherwise} \end{cases}$$

$$f_{X_1 | X_2 = x_2}(x_1) = \frac{f_{X_1, X_2}(x_1, x_2)}{f_{X_2}(x_2)}$$

$$= \begin{cases} \frac{2}{2(1-x_2)} & \text{if } 0 \leq x_1 \leq 1-x_2, \\ 0 & \text{otherwise} \end{cases}$$



$$= \begin{cases} \frac{1}{1-x_2} & \text{if } 0 \leq x_2 \leq 1-x_2, \\ 0 & \text{otherwise.} \end{cases}$$

Note that if  $x_2 = x_2$ , then  $Y = X_1 + X_2$ . We will now find  $f_{Y|X_2=x_2}$  by the method of transformations.

$$Y = g(x_2) = X_1 + x_2$$

Step 1:

$$(\alpha, \beta) = (0 + x_2, 1 - x_2 + x_2) = (x_2, 1)$$

Step 2:

$$X_1 = Y - x_2 = h(Y)$$

Step 3:

$$\frac{d}{dy} h(y) = 1 \neq 0 \text{ for } y \in (x_2, 1)$$

Step 4:

$$f_{Y|X_2=x_2}(y) = \begin{cases} f_{X_1}(h(y)) \left| \frac{d}{dy} h(y) \right| & \text{if } y \in (x_2, 1) \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} \frac{1}{1-x_2} & \text{if } y \in (x_2, 1), \\ 0 & \text{otherwise.} \end{cases}$$

Now,

$$f_Y(y) = \int_{-\infty}^{\infty} f_{Y|X_2=x_2}(y) f_{X_2}(x_2) dx_2$$



Note that  $f_{Y|X_2=x_2}(y) = 0$  if  $y < x_2$  or  $y > 1$ .

Hence, if  $0 \leq y \leq 1$ ,

$$\begin{aligned}
 f_Y(y) &= \int_0^y \frac{1}{1-x_2} 2(1-x_2) dx_2 \\
 &= \int_0^y 2 dx_2 \\
 &= 2y.
 \end{aligned}$$

Clearly if  $y < 0$  or  $y > 1$ ,

$$f_Y(y) = 0.$$

$$\text{Hence } f_Y(y) = \begin{cases} 2y & \text{if } 0 \leq y \leq 1, \\ 0 & \text{otherwise.} \end{cases}$$

6.26

(a)  $P(X < 2, Y > 1) = P((X, Y) \in A)$ ,

where  $A = \{(x, y) : x < 2, y > 1\}$

~~Range of~~

$R = \text{Range of density } f_{X,Y} = \{(x, y) : 0 \leq y \leq x < \infty\}$

$A \cap R = \{(x, y) : x < 2, y > 1, 0 \leq y \leq x < \infty\}$   
 $= \{(x, y) : 1 \leq x < 2, 1 \leq y \leq x\}$

$P((X, Y) \in A) = \int_A \int f_{X,Y}(x, y) dx dy = \int_1^2 \int_1^x e^{-x} dy dx$