Entropic Central Limit Theorem for Rényi Entropy

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Abstract—We establish a central limit theorem for Rényi entropies when the Rényi parameters belong to (0,1) for a large class of random vectors. This complements a celebrated result of Barron (1986). As an application, we show that a general Rényi entropy power inequality fails when the Rényi parameter is in (0,1).

Index Terms—Rényi entropy; entropy power inequality; entropic central limit theorem

I. INTRODUCTION

The Central Limit Theorem (henceforth, CLT) is a fundamental result in probability theory and statistics. It also has a plethora of applications in applied sciences. The CLT has close connections with information theory via the Entropy Power Inequality (henceforth, EPI) of Shannon [25] and Stam [26].

Let X be a random vector in \mathbb{R}^d with density f. The Shannon differential entropy h(X) is defined as

$$h(X) = -\int_{\mathbb{R}^d} f(x) \log f(x) dx.$$
(1)

The entropy power N(X) is defined as

$$N(X) = e^{2h(X)/d}.$$
 (2)

Shannon-Stam's EPI states that

$$N(X+Y) \ge N(X) + N(Y) \tag{3}$$

holds for arbitrary independent random vectors X and Y in \mathbb{R}^d . Let $\{X_n\}_{n\geq 1}$ be a sequence of independent and identically distributed (henceforth, i.i.d.) centered random vectors in \mathbb{R}^d with finite covariance matrix. We denote by Z_n the normalized sum

$$Z_n = \frac{X_1 + \dots + X_n}{\sqrt{n}}.$$
(4)

By induction and the homogeneity of $N(\cdot)$, the EPI (3) implies that the dyadic sequence $\{h(Z_{2^n})\}_{n\geq 1}$ is non-decreasing. It is well known that Gaussian maximizes the Shannon entropy when the covariance matrix is fixed (see, e.g., [10]). Hence, we have $h(Z_{2^n}) \leq h(Z)$, where Z is a centered Gaussian vector with the same covariance matrix as X_1 . This implies the convergence of the dyadic sequence $\{h(Z_{2^n})\}_{n\geq 1}$. A celebrated result of Barron [2] states that

$$h(Z_n) \to h(Z),$$
 (5)

as $n \to +\infty$, which strengthens the classical CLT that the sequence $\{Z_n\}_{n\geq 1}$ converges to Z in distribution. The

entropic CLT (5) along with the monotonicity of $\{h(Z_{2^n})\}_{n\geq 1}$ and maximization of entropy by Gaussians can be seen as an analogue of the second law of thermodynamics (see, e.g., [13]). Furthermore, it was proven in [1] that the sequence $\{h(Z_n)\}_{n\geq 1}$ is actually monotone at each step, and not only along dyadic steps, see also [17].

There has been recent success in extending the Shannon EPI (3) to the Rényi setting (see [4], [5], [7], [14], [18], [19], [21], [23], [24]) but few results are known about CLT for r-Rényi entropy (see [8] for r > 1 in dimension 1; see also [6] for convergence in Rényi divergence, which is not equivalent to convergence in Rényi entropy unless r = 1). The aim of this paper is to extend the above Shannon entropic CLT to the Rényi entropic setting, with particular interest in application to Rényi EPI. For $r \ge 0$, the r-Rényi entropy of a random vector X in \mathbb{R}^d with density f is defined as

$$h_r(X) = \frac{1}{1-r} \log \int_{\mathbb{R}^d} f(x)^r dx.$$
 (6)

For $r \in \{0, 1, \infty\}$, the definition is understood in the limiting sense, and $h_1(X)$ is the Shannon differential entropy. The *r*-Rényi entropy power $N_r(X)$ is defined as

$$N_r(X) = e^{2h_r(X)/d}$$
. (7)

This paper is organized as follows. The next section is dedicated to the convergence of $h_r(Z_n)$. For r > 1 convergence is fully characterized for random vectors in \mathbb{R}^d , while for $r \in (0, 1)$ sufficient conditions with application to Rényi EPI are explored. Explicitly, convergence is proven for log-concave random vectors and random vectors with radially symmetric unimodal densities and compact support. The last section presents applications of the Rényi entropic CLT. Most significantly, we show that a general *r*-Rényi EPI fails when $r \in (0, 1)$. The reader is directed to the full paper expanding on this work [15].

II. CENTRAL LIMIT THEOREM FOR RÉNYI ENTROPY

A fundamental tool in establishing various central limit theorems is the characteristic function. Recall that the characteristic function of a random vector X in \mathbb{R}^d is defined by

$$\varphi_X(t) = \mathbb{E}[e^{i\langle t, X \rangle}], \quad t \in \mathbb{R}^d.$$
(8)

Before providing sufficient conditions for convergence in r-Rényi entropy, with $r \in (0, 1)$, we first extend [8, Theorem 1.1] to the higher dimensional setting.

Theorem 2.1: Let r > 1. Let X_1, \ldots, X_n be i.i.d. centered random vectors in \mathbb{R}^d . Let us define Z_n as in (4) and denote by ρ_n the density of Z_n . Then, the following statements are equivalent.

- h_r (Z_n) → h_r(Z), where Z is a Gaussian random vector with mean 0 and same covariance matrix as X₁.
- 2) $h_r(Z_{n_0})$ is finite for some integer n_0 .
- 3) $\int_{\mathbb{R}^d} |\varphi_{X_1}(t)|^{\nu} dt < +\infty$ for some $\nu \ge 1$.
- 4) Z_{n_0} has a bounded density ρ_{n_0} for some integer n_0 .
- *Proof:* $1 \Longrightarrow 2$: Assume that $h_r(Z_n) \to h_r(Z)$ as $n \to +\infty$. Then, there exists an integer n_0 such that

$$h_r(Z) - 1 < h_r(Z_{n_0}) < h_r(Z) + 1.$$
 (9)

Since $h_r(Z)$ is finite, we deduce that $h_r(Z_{n_0})$ is finite as well. $2 \implies 3$: Assume that $h_r(Z_{n_0})$ is finite for some integer n_0 . Then, Z_{n_0} has a density ρ_{n_0} which is in $L^r(\mathbb{R}^d)$, and thus Z_n has a density $\rho_n \in L^r$ for any $n \ge n_0$ by the convolution structure of Z_n .

Case 1: If $r \geq 2$, then $\rho_n \in L^2(\mathbb{R}^d)$. Hence by Plancherel identity, $\varphi_{Z_n} \in L^2(\mathbb{R}^d)$. It follows that

$$\int_{\mathbb{R}^d} |\varphi_{Z_n}(t)|^2 dt = \int_{\mathbb{R}^d} |\varphi_{X_1}\left(t/\sqrt{n}\right)|^{2n} dt < +\infty.$$
(10)

We deduce that for $\nu = 2n_0$,

$$\int_{\mathbb{R}^d} |\varphi_{X_1}(t)|^{\nu} \, dt < +\infty. \tag{11}$$

Case 2: If $r \in (1, 2)$, then by the Hausdorff-Young inequality,

$$\|\varphi_{Z_n}\|_{L^{r'}} \le \frac{1}{(2\pi)^{d/r'}} \|\rho_n\|_{L^r},$$
(12)

where r' is the conjugate of r. Hence, for $\nu = r'n_0$,

$$\int_{\mathbb{R}^d} |\varphi_{X_1}(t)|^{\nu} dt < +\infty.$$
(13)

 $3 \implies 4$: Since $\int_{\mathbb{R}^d} |\varphi_{X_1}(t)|^{\nu} dt < +\infty$ for some $\nu \ge 1$, one may apply Gnedenko's local limit theorems (see [11]), which is valid in arbitrary dimension (see [3]). In particular,

$$\lim_{n \to +\infty} \sup_{x \in \mathbb{R}^d} |\rho_n(x) - \phi_{\Sigma}(x)| = 0,$$
(14)

where ϕ_{Σ} denotes the density of a Gaussian random vector with mean 0 and same covariance matrix as X_1 . We deduce that there exists an integer n_0 and a constant M > 0 such that $\rho_n \leq M$ for all $n \geq n_0$.

 $4 \implies 1$: Since ρ_{n_0} is bounded, then $\rho_{n_0} \in L^2$, and we deduce by Plancherel identity that $\int_{\mathbb{R}^d} |\varphi_{X_1}(t)|^{\nu} dt < +\infty$ for $\nu = 2n_0$. Hence (14) holds and there exists M > 0 such that $\rho_n \leq M$ for all $n \geq n_0$. Let us show that $\int \rho_n^r \to \int \phi_{\Sigma}^r$ as $n \to +\infty$, where ϕ_{Σ} denotes the density of a Gaussian random vector with mean 0 and same covariance matrix as X_1 . From the central limit theorem, there exists T > 0 such that for all n large enough,

$$\int_{|x|>T} \rho_n(x) dx < \varepsilon, \qquad \int_{|x|>T} \phi_{\Sigma}(x) dx < \varepsilon.$$
(15)

Hence,

$$\int_{|x|>T} \rho_n^r(x) dx \le M^{r-1} \int_{|x|>T} \rho_n(x) dx < M^{r-1}\varepsilon, \quad (16)$$

and similarly for $\int_{|x|>T} \phi_{\Sigma}^r$. Hence, for any $\delta > 0$, there exists T > 0 such that for all n large enough,

$$\left| \int_{|x|>T} \rho_n^r(x) dx - \int_{|x|>T} \phi_{\Sigma}^r(x) dx \right| < \delta.$$
 (17)

On the other hand, by (14), for all T > 0, the function $\rho_n^r(x) \mathbb{1}_{\{|x| \le T\}}$ converges everywhere to $\phi_{\Sigma}^r(x) \mathbb{1}_{\{|x| \le T\}}$ as $n \to +\infty$. Since $\rho_n^r(x) \mathbb{1}_{\{|x| \le T\}}$ is dominated by the integrable function $M^r \mathbb{1}_{\{|x| \le T\}}$, one may use the Lebesgue dominated theorem to conclude that

$$\lim_{n \to +\infty} \left| \int_{|x| \le T} \rho_n^r(x) dx - \int_{|x| \le T} \phi_{\Sigma}^r(x) dx \right| = 0.$$
(18)

Remark 2.1: When $r \in (0, 1)$, the statement of Theorem 2.1 fails since it is possible to find a bounded density ρ such that $\int \rho(x)^r dx = +\infty$ (e.g., Cauchy-type distributions). In particular, for random variable X_1 with such density, the implication $4 \Longrightarrow 2$ (and thus $4 \Longrightarrow 1$) in Theorem 2.1 does not necessarily hold when $r \in (0, 1)$ since by Jensen inequality $h_r(Z_n) \ge h_r(X_1/\sqrt{n}) = \infty$, for all $n \ge 1$. It was noted by Barron [2] that in the Shannon entropy case r = 1, the implication $1 \Longrightarrow 4$ does not necessarily hold.

In the next results, we provide sufficient conditions for a CLT to hold for *r*-Rényi entropy, with $r \in (0, 1)$, for different classes of random vectors in \mathbb{R}^d . Recall the definition of Z_n in (4).

Theorem 2.2: Let $r \in (0,1)$. Let $\{X_n\}_{n\geq 1}$ be a sequence of i.i.d. centered log-concave random vectors in \mathbb{R}^d . Then, $h_r(Z_n) < +\infty$ for all $n \geq 1$, and

$$h_r(Z_n) \to h_r(Z),$$

where Z is a Gaussian random vector with mean 0 and same covariance matrix as X_1 .

Proof: Since log-concavity is preserved under independent sum [22], Z_n is log-concave for all $n \ge 1$. Hence, for all $n \ge 1$, Z_n has a bounded log-concave density ρ_n , which satisfies

$$\rho_n(x) \le e^{-a_n |x| + b_n},\tag{19}$$

for all $x \in \mathbb{R}^d$, for some constants $a_n > 0$, $b_n \in \mathbb{R}$ possibly depending on the dimension (see, e.g., [9]). Hence, for all $n \ge 1$,

$$\int_{\mathbb{R}^d} \rho_n(x)^r \, dx \le \int_{\mathbb{R}^d} e^{-r(a_n|x|+b_n)} \, dx < +\infty.$$
⁽²⁰⁾

We deduce that $h_r(Z_n) < +\infty$ for all $n \ge 1$.

The boundedness of ρ_n implies that (14) holds, and thus there exists an integer n_0 such that for all $n \ge n_0$,

$$\rho_n(0) > \frac{1}{2}\phi_{\Sigma}(0),\tag{21}$$

where Σ is the covariance matrix of X_1 (and thus does not depend on *n*). Moreover, since ρ_n is log-concave, one has for all $x \in \mathbb{R}^d$,

$$\rho_n(rx) = \rho_n((1-r)0 + rx)
\geq \rho_n(0)^{1-r}\rho_n(x)^r
\geq \frac{1}{2^{1-r}}\phi_{\Sigma}(0)^{1-r}\rho_n(x)^r.$$
(22)

Hence, for all T > 0,

$$\int_{|x|>T} \rho_n(x)^r \, dx \le \frac{2^{1-r}}{\phi_{\Sigma}(0)^{1-r}} \int_{|x|>T} \rho_n(rx) \, dx \qquad (23)$$

$$=\frac{2^{1-r}}{r^d \phi_{\Sigma}(0)^{1-r}} \mathbb{P}(|Z_n| > rT)$$
(24)

$$\leq \frac{1}{T^2} \frac{2^{1-r} \mathbb{E}[|X_1|^2]}{r^{d+2} \phi_{\Sigma}(0)^{1-r}},$$
(25)

where the last inequality follows from Markov's inequality and the fact that

$$\mathbb{E}[|Z_n|^2] = \frac{\mathbb{E}[|X_1|^2] + \dots + \mathbb{E}[|X_n|^2]}{n} = \mathbb{E}[|X_1|^2].$$
 (26)

Hence, for every $\varepsilon > 0$, one may choose a positive number T such that for all n large enough,

$$\int_{|x|>T} \rho_n^r(x) dx < \varepsilon, \qquad \int_{|x|>T} \phi_{\Sigma}^r(x) dx < \varepsilon, \qquad (27)$$

and hence

$$\left| \int_{|x|>T} \rho_n^r(x) dx - \int_{|x|>T} \phi_{\Sigma}^r(x) dx \right| < \varepsilon.$$
 (28)

On the other hand, from (14), we conclude as in the proof of Theorem 2.1 that for all T > 0,

$$\lim_{n \to +\infty} \left| \int_{|x| \le T} \rho_n^r(x) dx - \int_{|x| \le T} \phi_{\Sigma}^r(x) dx \right| = 0.$$
 (29)

Next, we provide a convergence result for the more general class of unimodal distributions under additional symmetry assumptions. First, we need the following stability result.

Proposition 2.1: The class of unimodal spherically symmetric random variables is stable under convolution.

Proof: Suppose that f_i are densities such that $f_i(Tx) = f_i(x)$ for an orthogonal map T and $|x| \le |y|$ implies $f_i(x) \ge f_i(y)$. By the layer cake decomposition, we write

$$f_i(x) = \int_0^\infty \mathbb{1}_{\{(u,v):f_i(u)>v\}}(x,\lambda)d\lambda.$$
 (30)

After applying Fubini-Tonelli,

$$\begin{aligned}
f_1 \star f_2(x) & (31) \\
\int & f_1(x) + f_2(x) \\
\end{bmatrix}$$

$$= \int_{\mathbb{R}^d} f_1(x-y) f_2(y) dy$$

= $\int_0^\infty \int_0^\infty \left(\int_{\mathbb{R}^d} \mathbb{1}_{\{(u,v): f_1(u) > v\}}(x-y,\lambda_1) \right)$ (32)

$$\times \mathbb{1}_{\{(u,v):f_2(u)>v\}}(y,\lambda_2)dy \bigg) d\lambda_1 d\lambda_2.$$
(33)

Notice that by the spherical symmetry and decreasingness of f_i , the super-level set

$$L_{\lambda_i} = \{ u : f_i(u) > \lambda_i \}$$
(34)

is an origin symmetric ball. Thus we can write the integrand in (33) as

$$\int_{\mathbb{R}^d} \mathbb{1}_{L_{\lambda_1}}(x-y) \mathbb{1}_{L_{\lambda_2}}(y) dy = \mathbb{1}_{L_{\lambda_1}} \star \mathbb{1}_{L_{\lambda_2}}(x).$$
(35)

This quantity is clearly dependent only on |x|, giving spherical symmetry. Additionally as the convolution of two log-concave functions, $\mathbb{1}_{L_{\lambda_1}} \star \mathbb{1}_{L_{\lambda_2}}$ is log-concave as well. It follows that for every λ_1, λ_2 , and $|x| \leq |y|$ we have

$$\mathbb{1}_{L_{\lambda_1}} \star \mathbb{1}_{L_{\lambda_2}}(x) \ge \mathbb{1}_{L_{\lambda_1}} \star \mathbb{1}_{L_{\lambda_2}}(y). \tag{36}$$

Then we can finish the proof by integrating both sides of the above inequality.

Let us establish large deviation and pointwise inequalities for radially symmetric unimodal densities with compact support.

Theorem 2.3 (Hoeffding [12]): For independent random variables X_i with zero mean, bounded in $(a_i, b_i), i = 1, ..., n$, one has for all T > 0,

$$\mathbb{P}\left(\sum_{i=1}^{n} X_i > T\right) \le \exp\left\{-\frac{2T^2}{\sum_{i=1}^{n} (b_i - a_i)^2}\right\}.$$
 (37)

Lemma 2.1: For centered independent random vectors X_i in \mathbb{R}^d satisfying $\mathbb{P}(|X_i| > R) = 0$, $i = 1, \ldots, n$, for some R > 0, one has for all T > 0,

$$\mathbb{P}\left(\left|\frac{X_1 + \dots + X_n}{\sqrt{n}}\right| > T\right) \le 2d \exp\left\{-\frac{T^2}{2d^2R^2}\right\}.$$
 (38)

Proof: Let $X_{i,j}$ be the *j*-th coordinate of the random vector X_i . Then we have

$$\mathbb{P}\left(\left|\frac{X_1 + \dots + X_n}{\sqrt{n}}\right| > T\right) \tag{39}$$

$$\leq \mathbb{P}\left(\bigcup_{j=1}^{d} \left\{ |X_{1,j} + \dots + X_{n,j}| > \frac{T\sqrt{n}}{d} \right\} \right)$$
(40)

$$\leq \sum_{j=1}^{d} \mathbb{P}\left(|X_{1,j} + \dots + X_{n,j}| > \frac{T\sqrt{n}}{d} \right)$$
(41)

$$\leq 2d \exp\left(-\frac{T^2}{2d^2R^2}\right),\tag{42}$$

where inequality (40) follows from the pigeon-hole principle, (41) from a union bound, and (42) follows from applying Hoeffding's inequality (Theorem 2.3) to $X_{1,j} + \cdots + X_{n,j}$ and $(-X_{1,j}) + \cdots + (-X_{n,j})$.

We deduce the following pointwise estimate for unimodal radially symmetric and bounded random variables.

Corollary 2.1: If X_i are i.i.d. with radially symmetric unimodal density function supported on the Euclidean ball $B_R = \{x : |x| \le R\}$ for some R > 0, then letting ρ_n denote density of the normalized sum $(X_1 + \cdots + X_n)/\sqrt{n}$, there exists $c_d > 0$ such that for |x| > 2,

$$\rho_n(x) \le c_d \exp\left\{-\frac{(|x|-1)^2}{2d^2R^2}\right\}.$$
(43)

Proof: Stating Lemma 2.1 in terms of ρ_n , we have

$$\int_{|y|>T} \rho_n(w) dw \le 2d \exp\left\{-\frac{T^2}{2d^2 R^2}\right\}.$$
 (44)

Since the class of radially symmetric unimodal random variables is stable under independent summation by Proposition 2.1, ρ_n is radially symmetric and unimodal, so that

$$\rho_n(x) \le \frac{\int_{B_{|x|} \setminus B_{|x|-1}} \rho_n(w) dw}{\operatorname{Vol}(B_{|x|} \setminus B_{|x|-1})} \tag{45}$$

$$\leq \frac{\int_{|y|\geq |x|-1} \rho_n(w) dw}{(2^d-1)\omega_d} \tag{46}$$

where ω_d is the volume of the unit ball. Note that

$$\operatorname{Vol}(B_{|x|} \setminus B_{|x|-1}) = (|x|^d - (|x|-1)^d)\omega_d \ge (2^d - 1)\omega_d,$$
(47)

since $t \mapsto t^d - (t-1)^d$ is increasing, so that (46) follows. Now applying (44) we have

$$\rho_n(x) \le \frac{\int_{|y|\ge|x|-1} \rho_n(w) dw}{(2^d - 1)\omega_d} \tag{48}$$

$$\leq \frac{2d}{(2^d - 1)\omega_d} \exp\left\{-\frac{(|x| - 1)^2}{2d^2 R^2}\right\}$$
(49)

and our result holds with

$$c_d = \frac{2d}{(2^d - 1)\omega_d}.$$
 (50)

We are now ready to establish a convergence result for unimodal radially symmetric bounded random vectors.

Theorem 2.4: Let $r \in (0, 1)$. Let $\{X_n\}_{n \ge 1}$ be a sequence of i.i.d. random vectors in \mathbb{R}^d with a radially symmetric unimodal density with compact support. Then,

$$\lim_{n \to \infty} N_r \left(\frac{X_1 + \dots + X_n}{\sqrt{n}} \right) = N_r(Z)$$
 (51)

where Z is a Gaussian random vector with mean 0 and same covariance matrix as X_1 .

Proof: Let us denote by ρ_n the density of the normalized sum $(X_1 + \cdots + X_n)/\sqrt{n}$. Since $\rho_1 \in L^1$, it follows that ρ_n , $n \ge 2$, are continuous, and since ρ_n are, in addition, radially symmetric unimodal densities by Proposition 2.1, then ρ_n , $n \ge 2$, are bounded. Hence, the point 3) of Theorem 2.1 holds with $\nu = 2$, and one may thus apply (14) together with Lebesgue dominated convergence to conclude that for all T > 0,

$$\lim_{n \to +\infty} \left| \int_{|x| \le T} \rho_n^r(x) dx - \int_{|x| \le T} \phi_{\Sigma}^r(x) dx \right| = 0.$$
 (52)

On the other hand, by Corollary 2.1, for all $\varepsilon > 0$, one may choose T > 0 such that for all $n \ge 1$,

$$\int_{|x|>T} \rho_n^r(x) dx < \varepsilon, \qquad \int_{|x|>T} \phi_{\Sigma}^r(x) dx < \varepsilon, \qquad (53)$$

and hence

$$\left| \int_{|x|>T} \rho_n^r(x) dx - \int_{|x|>T} \phi_{\Sigma}^r(x) dx \right| < \varepsilon.$$
 (54)

III. NO RÉNYI EPI OF ORDER $r \in (0, 1)$

Definition 3.1: For $r \in [0, \infty]$, define c_r as the largest number such that for any finite sequence of independent random vectors X_i in \mathbb{R}^d , the inequality

$$N_r(X_1 + \dots + X_n) \ge c_r \sum_{i=1}^n N_r(X_i)$$
 (55)

holds.

Theorem 3.1: For $r \in (0, 1)$, the constant c_r defined in (55) satisfies

$$c_r = 0. (56)$$

Theorem 3.1 affirms a striking difference between Rényi EPI of parameter $r \ge 1$ and $r \in (0, 1)$. Indeed, it was shown in [5] that for $r \ge 1$, one has

$$c_r \ge \frac{1}{e} r^{\frac{1}{r-1}}.$$
(57)

Theorem 3.1 can be reformulated as follows.

Theorem 3.2: For any $r \in (0,1)$ and $\varepsilon > 0$, there exist independent random vectors X_1, \ldots, X_n in \mathbb{R}^d , for some $d \ge 1$ and $n \ge 2$, such that

$$N_r(X_1 + \dots + X_n) < \varepsilon \sum_{i=1}^n N_r(X_i).$$
(58)

The motivating observation for this line of argument is the fact that for $r \in (0, 1)$, there exist variables with finite variance and infinite r-Rényi entropy. One might anticipate that this could contradict the existence of an r-Rényi EPI, as the central limit theorem forces i.i.d. sums to become "more Gaussian". Heuristically, one anticipates for large n, and X_i drawn from such a distribution, that $N_r(X_1 + \cdots + X_n)/n = N_r((X_1 + \cdots + X_n)/\sqrt{n})$ should approach $N_r(Z)$, where Z is a Gaussian with variance equal to X_i 's, while $\sum N_r(X_i)/n = N_r(X_1)$ is infinite.

Proof of Theorem 3.1: Let us consider the following density

$$f(x) = f_{R,p,d}(x) = C_R (1+|x|)^{-p} \mathbb{1}_{B_R}(x), \quad x \in \mathbb{R}^d,$$
(59)

with p, R > 0 and C_R implicitly determined to make f a density. Note that f is bounded, unimodal, and radially symmetric. Thus its covariance matrix is a multiple of the identity, i.e., $\sigma_R^2 I$ for some $\sigma_R > 0$. Computing in spherical coordinates one can easily see that $\lim_{R\to\infty} C_R$ is finite for p > d, and we can thus define a density $f_{\infty,p,d}$. What is

more, when p > d+2, the limiting density $f_{\infty,p,d}$ has a finite covariance matrix, and has finite Rényi entropy if and only if p > d/r.

Now fix $r \in (0,1)$ and take the dimension to be $d^* = \min\{d \in \mathbb{N} : d > 2r/(1-r)\}$, and $p \in (d^* + 2, d^*/r)$. In this case, the limit density f_{∞,p,d^*} is well defined and it has finite covariance matrix $\sigma_{\infty}^2 I$, but the corresponding *r*-Rényi entropy is infinite. Now we select independent random vectors X_1, \dots, X_n from the distribution f_{R,p,d^*} . Since f_{R,p,d^*} is a radially symmetric unimodal density with compact support, one may apply Theorem 2.4 to conclude that

$$\lim_{n \to \infty} N_r \left(\frac{X_1 + \dots + X_n}{\sqrt{n}} \right) = \sigma_R^2 N_r(Z_{Id}), \quad (60)$$

where Z_{Id} is the standard *d*-dimensional Gaussian. Notice that

$$\lim_{R \to \infty} N_r(X_1) = \infty, \tag{61}$$

while

$$\lim_{R \to \infty} \sigma_R = \sigma_\infty < \infty.$$
 (62)

Given M > 0, we can take R large enough such that $N_r(X_1) \ge M$, and $|\sigma_R^2 - \sigma_\infty^2| \le 1$. Then we can take n large enough such that

$$N_r\left(\frac{X_1 + \dots + X_n}{\sqrt{n}}\right) \le (\sigma_{\infty}^2 + 2)N_r(Z_{Id}).$$
(63)

We conclude that for the inequality (55) to hold. We must have

$$c_r \le \frac{(\sigma_\infty^2 + 2)N_r(Z_{Id})}{M} \tag{64}$$

for all M > 0. Taking $M \to \infty$ this can only hold if $c_r = 0$.

Remark 3.1: The counterexample we have built in the proof of Theorem 3.1 is a random vector in \mathbb{R}^d with an *s*-concave density with s < 0 and $|s| > \frac{r}{d}$. In [16], where the terminology of *s*-concave density is recalled, we prove that for $|s| < \frac{r}{d}$, a Rényi EPI does hold for fixed $r \in (0, 1)$, hence complementing the above negative result and extending the following first announced in [20].

Theorem 3.3 ([21]): For $r \in (0, 1)$, there exists C(r) > 0 such that for all X_1, \ldots, X_n independent log-concave random vectors in \mathbb{R}^d ,

$$N_r(X_1 + \dots + X_n) \ge C(r) \sum_{i=1}^n N_r(X_i).$$
 (65)

In particular one can take $C(r) = e r^{\frac{1}{1-r}}$.

By applying Theorem 2.2 to X_i distributed according to a Laplace distribution, which is log-concave, it follows that the optimal value C(r) satisfying (65) for all log-concave random vectors verifies $C(r) \leq \pi r^{\frac{1}{1-r}}$.

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